

lm.br: An R Package for Broken Line Regression

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Abstract

The R package `lm.br` delivers exact tests and exact confidence regions for a changepoint in linear or multivariate linear regression. This package implements the likelihood theory of conditional inference. Examples demonstrate its use and note some properties of the broken line models.

1 Theory

A broken-line model consists of two straight lines joined at a changepoint. Three variants are

$$y_i = \alpha + \beta(x_i - \theta)_- + \beta'(x_i - \theta)_+ + e_i \quad (1)$$

$$y_i = \alpha + \beta(x_i - \theta)_- + e_i \quad (2)$$

$$y_i = \beta(x_i - \theta)_- + e_i \quad (3)$$

where $e \sim N(0, \sigma^2 \Sigma)$, denoting $a_- = \min(a, 0)$ and $a_+ = \max(a, 0)$. Parameters θ , α , β , β' , σ are unknown but Σ is known. Model (2) is a threshold model, while model (3) would apply for a known threshold level. The following presentation orders $x_1 \leq x_2 \leq \dots \leq x_n$, without loss of generality.

The likelihood-ratio is a test statistic. A test statistic D assigns a numeric value to a postulate parameter value, p_0 , based on the model and the observations. $D(p_0)$ is itself a random variable because it is a function of the random observations. A significance level is the probability that D could be worse than the observed value, $SL(p_0) = Pr[D(p_0) > D(p_0)_{obs}]$, based on the model. The set of postulate values such that $SL > \alpha$ is a $100(1 - \alpha)\%$ confidence region.

Conditional inference incorporates sufficient statistics to account for the other, unknown parameters. This refinement determines the exact distribution of a test statistic, even for small data sets. Student's t , for example, is the distribution of a sample mean conditional on a sufficient statistic for the variance. See Kalbfleisch (1985, ch.15).

Knowles, Siegmund, and Zhang (1991) derived the conditional likelihood-ratio (CLR) significance tests for the non-linear parameter in semilinear regression. Siegmund and Zhang (1994) applied these tests to get exact confidence intervals

for the changepoint θ in models (1) and (2), and exact confidence regions for the two-parameter changepoint (θ, α) in model (2). Knowles et al. (1991) also developed a formula for rapid evaluation, which `lm.br` implements.

`lm.br` extends this theory. Their method also derives an exact significance test for (θ, α) in model (1). The theory adapts to the case σ known. And these exact significance tests simplify for a postulate changepoint value outside of $[x_1, x_n]$ (Knowles and Siegmund, 1989).

Approximate-F (AF) is another inference method that is common in nonlinear regression, but it is not exact. The AF method estimates the distribution of a likelihood-ratio statistic by its asymptotic χ^2 distribution, with partial conditioning on a sufficient statistic for the variance. See Draper and Smith (1998, sec. 24.6).

2 Examples

2.1 Simulation Tests

Table: Coverage frequencies of the 95% confidence interval for 100 random models

		CLR	AF
10 observations,	$x_1 - 1 < \theta < x_n + 1$	95.0 – 95.2	90.0 – 97.5
30 observations,	$x_{10} < \theta < x_{20}$	95.0 – 95.2	90.8 – 95.0
100 observations,	$x_{10} < \theta < x_{20}$	95.0 – 95.2	91.3 – 95.0

To give one specific example, coverage frequency is 95.2% by CLR but 90.7% by AF for a first-line slope -1, second-line slope +0.5, changepoint $\theta = 3$, and 10 observations at $x = (1.0, 1.1, 1.3, 1.7, 2.4, 3.9, 5.7, 7.6, 8.4, 8.6)$ with $\sigma = 1$.

The formulae that generated the random models are

$$\begin{aligned}
 n = 10 \quad x_1 = 1 \quad x_i = x_{i-1} + 2U \text{ for } i > 1 \quad \theta = x_1 - 1 + (x_n - x_1 + 2)U \\
 \alpha = 0 \quad \beta = -1 \quad \beta' = 2 - 2.5U \quad \sigma = 0.1 + 2U
 \end{aligned}$$

or $n = 30$ or $n = 100$ and $\theta = x_{10} + (x_{20} - x_{10})U$, where $U \sim \text{Uniform}(0, 1)$. For each model, the program output one million sets of random $y_i = \alpha + \beta(x_i - \theta)_- + \beta'(x_i - \theta)_+ + \sigma N(0, 1)$ and counted how often $SL(\theta) > .05$. Coverage frequencies should be accurate to $\pm 0.05\%$.

2.2 Broken Line Regression

Drinking and driving might have followed a broken-line trend. Yearly surveys were adjusted by a seasonal index based on monthly surveys for a similar question (TIRF, 1998–2007; CAMH, 2003). The annual surveys asked respondents if in the past 30 days they had driven within two hours after a drink, while the monthly surveys asked if in the past 30 days they had driven within one hour after two drinks. Figure 1 shows the survey results without and with seasonal adjustment, and the exact 90% confidence region for the changepoint if the adjustment were valid.

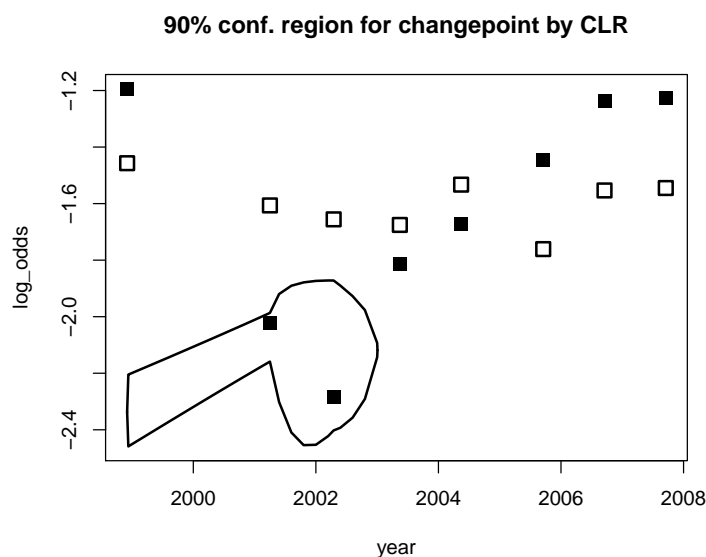


Figure 1: Drinking-and-driving surveys log-odds (blank squares) and log-odds with seasonal adjustment (solid squares) versus year, and the exact 90% confidence region for the changepoint (θ, α) .

In R the commands are

```
> library( lm.br )
> log_odds <- c( -1.194, -2.023, -2.285, -1.815, -1.673,
+   -1.444, -1.237, -1.228 )
> year <- c( 1998.92, 2001.25, 2002.29, 2003.37, 2004.37,
+   2005.71, 2006.71, 2007.71 )
> VarCov <- matrix( c( 0.0361, 0, 0, 0, 0, 0, 0, 0,
+   0, 0.0218, 0.0129, 0, 0, 0, 0, 0,
+   0, 0.0129, 0.0319, 0, 0, 0, 0, 0,
```

```

+ 0, 0, 0, 0.0451, 0.0389, 0, 0, 0,
+ 0, 0, 0, 0.0389, 0.0445, 0, 0, 0,
+ 0, 0, 0, 0, 0, 0.0672, 0.0607, 0.0607,
+ 0, 0, 0, 0, 0, 0.0607, 0.0664, 0.0607,
+ 0, 0, 0, 0, 0, 0.0607, 0.0607, 0.0662 ),
+ nrow = 8, ncol = 8 )
> dd <- lm.br( log_odds ~ year, w= VarCov, inv= T, var.k= T )
> dd$ci( )

```

```

95-percent confidence interval for changepoint 'theta' by CLR
[ 2001.29, 2002.88 ]

```

```

> dd$ci( method = "AF" )

```

```

95-percent confidence interval for changepoint 'theta' by AF
[ 1998.92, 2002.82 ]

```

The wide difference between the CLR and AF confidence intervals here is due to plateaus in the significance levels on end-intervals. Both the CLR and AF methods give a constant significance level for all postulate values θ_0 on $[x_1, x_2]$, on $[x_{n-1}, x_n]$, and outside $[x_1, x_n]$, in model (1). The inference assumes that any line slope is possible, extending to an instantaneous drop near Dec. 1998 in this example.

2.3 Multivariate Regression

`lm.br` can test for a changepoint in multivariate linear regression. `lm.br` tests for a coefficient change in the first term of the regression model, assuming continuity. It does not test for an arbitrary structural change that might involve multiple parameters or discontinuity.

Liu, Wu, and Zidek (1997) suggested a changepoint for the coefficient of car weight in a linear fit of miles-per-gallon against weight and horsepower for 38 cars, 1978-79 models. One of R's included datasets is the ratings for 32 cars, 1973-74 models. Analysis of this 1973-74 dataset by conditional likelihood-ratio inference also shows some evidence for a changepoint:

```

> lm.br( mpg ~ wt + hp, data = mtcars )

```

Call:

```

lm.br(formula = mpg ~ wt + hp, data = mtcars)

```

Broken-line type: LL

```

Significance Level of H0:"no changepoint" vs H1:"one changepoint"
SL= 0.0110841 for theta0 = 1.32 by method CLR

```

95-percent confidence interval for changepoint 'theta' by CLR
 [2.13813, 5.14625]

Changepoint and coefficients:

theta	1-vector	wt < theta	wt > theta	hp
2.62000	25.02750	-8.81519	-2.51738	-0.03003

`lm.br` applies a canonical transform for multivariate regression (Siegmund and Zhang, 1994). One way to see how this method works is formulaic. The composite likelihood-ratio statistic uses optimal values for unknown parameters. A canonical model lets these optimal coefficients for other terms reduce their correspondent errors to zero always. Thus they have no effect on inference, so the algebra can omit them. This elimination reduces the multivariate model to a univariate model. See Hoffman and Kunze (1971, ch.6), Lehmann and Romano (2005, sec. 7.1).

3 Remarks

If a broken line with Normal errors represents the relationship between a factor and responses, `lm.br` solves the inference step for the changepoint. Fitting a broken line can reveal the plausible interval for a changepoint, although practical cause-effect relations usually have a smooth transition. Any statistical analysis should examine the fit of the model and the error distribution with graphs and significance tests, interpret results, and consider adjustments to the model or alternative models.

References

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