



R Functions to Symbolically Compute the Central Moments of the Multivariate Normal Distribution

Kem Phillips
Sourland Biostatistics

Abstract

The central moments of the multivariate normal distribution are functions of its $n \times n$ variance-covariance matrix Σ . These moments can be expressed symbolically as linear combinations of products of powers of the elements of Σ . A formula for these moments derived by differentiating the characteristic function is developed. The formula requires searching integer matrices for matrices whose n successive row and column sums equal the n exponents of the moment. This formula is implemented in R, with R functions to display moments in L^AT_EX and to evaluate moments at specified variance-covariance matrices included.

Keywords: central moments, multivariate normal distribution, symbolic computation, R, L^AT_EX.

1. Introduction

The central moments of an n -dimensional random vector X are defined as

$$m_{k_1, \dots, k_n} = E[(X_1 - \mu_1)^{k_1} (X_2 - \mu_2)^{k_2} \cdots (X_n - \mu_n)^{k_n}], \quad (1)$$

where $E[\cdots]$ denotes expected value. Suppose that X is distributed according to the multivariate normal distribution with mean μ and variance-covariance matrix

$$\Sigma = (\sigma_{ij}) \quad (2)$$

where the variance terms are $\sigma_{ii}, i = 1, \dots, n$, the covariance terms are $\sigma_{ij}, i \neq j$, and by symmetry $\sigma_{ij} = \sigma_{ji}$. For the multivariate normal distribution the central moments are not functions of the mean vector μ , and depend only on the variance-covariance terms σ_{ij} .

Simple cases are familiar. Setting μ to 0,

$$m_2 = E[X_1^2] = \sigma_{1,1} \quad (3)$$

2 Symbolic Computation of the Central Moments of the Multivariate Normal Distribution

$$m_{1,2} = E[X_1 X_2] = \sigma_{1,2} \quad (4)$$

Slightly more complicated cases can be computed directly or by manipulating simple expressions obtained for moments of the form $E[X_1 \cdots X_n]$ (Wikipedia 2009). For example,

$$m_{2,2} = E[X_1^2 X_2^2] = 2\sigma_{1,2}^2 + \sigma_{1,1}\sigma_{2,2} \quad (5)$$

$$m_{2,1,1} = E[X_1^2 X_2^1 X_3^1] = 2\sigma_{1,2}\sigma_{1,3} + \sigma_{1,1}\sigma_{2,3} \quad (6)$$

$$m_{1,1} = E[X_1^3 X_2^1] = 3\sigma_{1,1}\sigma_{1,2} \quad (7)$$

Although some higher order moments follow known patterns, most are much harder to determine by simple calculations.

We wish to compute the symbolic expression of any moment m_{k_1, \dots, k_n} in terms of the $n(n+1)$ symbols σ_{ij} . Note that if $\sum_{i=1}^n k_i$ is an odd integer, the moment is 0, so that only moments where this sum is even need be considered.

The multivariate normal distribution is fundamental to mathematical statistics, and its moments play a central role in statistical methodology. Various methods have been developed to numerically compute them (Muirhead 1982, p. 46) and (Anderson 1971, p. 49). Kan (2008) developed a formula (his Proposition 1) for the central moments as a repeated sum. He gives an excellent review of other formulas that have been developed, and cites Isserlis (1918) as deriving the first expression for the central moments. Muirhead (p. 49) used the matrix derivatives of the multivariate normal distribution's characteristic function to derive a formula for multivariate cumulants. Tracy and Sultan (1993) also used matrix derivatives to derive an expression for the distribution's moments (their Theorem 2) based on a recurrence relationship of the derivatives. This article develops a new explicit formula for the moments starting with the derivatives of the characteristic function. The expression for the moments is based on a search algorithm over certain integer matrices. The final goal of this paper is to translate this formula into R functions that produce symbolic representations of moments in terms of the variance-covariance terms σ_{ij} .

The functions described here are available in the package **symmoments** implemented in the R system for statistical computing (R Development Core Team 2009). Both R itself and the **symmoments** package (as well as all other packages used in this paper) are available under the terms of the General Public License (GPL) from the Comprehensive R Archive Network (CRAN, <http://CRAN.R-project.org/>).

2. Development of the formula

The moments of any distribution can be represented by the derivatives of the distribution's characteristic function. The characteristic function of the multivariate normal distribution is (Muirhead 1982, p. 5, 49)

$$E[e^{it^\top X}] = e^{it^\top \mu - \frac{1}{2}t^\top \Sigma t} \quad (8)$$

where $t = (t_1, t_2, \dots, t_n)$. Within a constant, the moment is the k_1, \dots, k_n -order derivative of the characteristic function evaluated at $t = 0$:

$$m_{k_1, \dots, k_n} = i^{-\sum_{i=1}^n k_i} \frac{d^{\sum_{i=1}^n k_i}}{d^{k_1} t_1 d^{k_2} t_2 \cdots d^{k_n} t_n} E[e^{it^\top X}]|_{t=0} \quad (9)$$

where i is the imaginary unit. Expanding the exponential into an infinite sum, this is

$$m_{k_1, \dots, k_n} = i^{-\sum_{i=1}^n k_i} \frac{d^{\sum_{i=1}^n k_i}}{d^{k_1} t_1 d^{k_2} t_2 \dots d^{k_n} t_n} \sum_{\ell=0}^{\infty} (it^\top \mu - \frac{1}{2}(t^\top \Sigma t))^\ell / \ell! \big|_{t=0} \quad (10)$$

Since we are to compute the central moment, we will set $\mu = 0$, so that the term $it^\top \mu$ will not appear:

$$m_{k_1, \dots, k_n} = i^{-\sum_{i=1}^n k_i} \frac{d^{\sum_{i=1}^n k_i}}{d^{k_1} t_1 d^{k_2} t_2 \dots d^{k_n} t_n} \sum_{\ell=0}^{\infty} (-\frac{1}{2}(t^\top \Sigma t))^\ell / \ell! \big|_{t=0} \quad (11)$$

Since $\sum_{i=1}^n k_i$ is even, the term $i^{-\sum_{i=1}^n k_i} = (-1)^{-\sum_{i=1}^n k_i/2}$. We note that the term $i^{-\sum_{i=1}^n k_i}$ will ultimately cancel with the negative in the infinite sum, and will be omitted for convenience in notation.

The expression in t is

$$t^\top \Sigma t = \sum_{ij} \sigma_{ij} t_i t_j \quad (12)$$

We need to find the coefficient of $t_1^{\ell_1} t_2^{\ell_2} \dots t_n^{\ell_n}$ in

$$(t^\top \Sigma t)^\ell = \left(\sum_{ij} \sigma_{ij} t_i t_j \right)^\ell = \sum_{ij} \sigma_{ij} t_i t_j \dots \sum_{ij} \sigma_{ij} t_i t_j \quad (13)$$

All products in the elements of the sum will occur. Any product will be obtained by choosing $\sigma_{ij} t_i t_j$ a certain number of times, say ℓ_{ij} . Since one term is chosen from each of ℓ terms, $\sum_{ij} \ell_{ij} = \ell$. Further, for any such matrix (ℓ_{ij}) , there will be a term, since it can be constructed by choosing $\sigma_{ij} t_i t_j$ from the first ℓ_{ij} terms, and so forth for each (ij) until ℓ is exhausted. For any (ℓ_{ij}) there are

$$\binom{\ell}{\ell_{11} \dots \ell_{nn}} \quad (14)$$

ways to choose the terms, where this is the multinomial coefficient. So,

$$\left(\sum_{ij} \sigma_{ij} t_i t_j \right)^\ell = \sum_{\{(\ell_{ij}) | \sum_{ij} \ell_{ij} = \ell\}} \binom{\ell}{\ell_{11} \dots \ell_{nn}} \prod_{ij} (\sigma_{ij} t_i t_j)^{\ell_{ij}} \quad (15)$$

$$= \sum_{\{(\ell_{ij}) | \sum_{ij} \ell_{ij} = \ell\}} \binom{\ell}{\ell_{11} \dots \ell_{nn}} \prod_{ij} \sigma_{ij}^{\ell_{ij}} \prod_{ij} (t_i t_j)^{\ell_{ij}} \quad (16)$$

For the moment, we distinguish between σ_{ij} and σ_{ji} as symbols. As a result, each $\prod_{ij} \sigma_{ij}^{\ell_{ij}}$ is unique as determined by unique (ℓ_{ij}) . However, $t_i t_j = t_j t_i$, so since each σ_{ij} is combined with two t 's, the total exponent in t is 2ℓ . That is, a term $t_1^{\ell_1} t_2^{\ell_2} \dots t_n^{\ell_n}$ must have $\sum_{i=1}^n \ell_i = 2\ell$. We need to determine the terms for which, for any (ℓ_1, \dots, ℓ_n) ,

$$\prod_{ij} (t_i t_j)^{\ell_{ij}} = t_1^{\ell_1} \dots t_n^{\ell_n} \quad (17)$$

4 Symbolic Computation of the Central Moments of the Multivariate Normal Distribution

We will get t_k in the product in the following mutually exclusive cases:

Condition	Exponent of t_k
$i = k, j \neq k$	1
$i \neq k, j = k$	1
$i = j = k$	2

(18)

So the exponent of t_k will be

$$\sum_{i=k, j \neq k} \ell_{ij} + \sum_{i \neq k, j=k} \ell_{ij} + 2\ell_{kk} \quad (19)$$

This sum is obtained by adding the sum of ℓ_{ij} across row k to the sum across column k , since the diagonal element k occurs in both sums. That is, we get

$$\sum_i \ell_{ik} + \sum_j \ell_{kj} = \sum_i (\ell_{ik} + \ell_{ki}) = \ell_k \quad (20)$$

We can now partition the set of (ℓ_{ij}) in Equation 16 according to these sums, that is, $\{\ell_k, k = 1 \dots n\}$. As stated before, the sum of the exponents, ℓ_k , must be 2ℓ .

$$\left(\sum_{ij} \sigma_{ij} t_i t_j \right)^\ell = \sum_{\{(\ell_1, \dots, \ell_n) | \sum_k \ell_k = 2\ell\}} \sum_{\{(\ell_{11}, \dots, \ell_{nn}) | \sum_i (\ell_{ik} + \ell_{ki}) = \ell_k, k=1 \dots n\}} \left(\binom{\ell}{\ell_{11} \dots \ell_{nn}} \prod_{ij} \sigma_{ij}^{\ell_{ij}} \right) \prod_{i=1}^n t_i^{\ell_i} \quad (21)$$

Since differentiation is distributive with respect to addition and multiplication by constants, the derivative of the product of t s can be determined from the derivatives of the individual terms:

$$\frac{d^{\sum_{i=1}^n k_i}}{d^{k_1} t_1 d^{k_2} t_2 \dots d^{k_n} t_n} \prod_{i=1}^n t_i^{\ell_i} = \prod_{i=1}^n \frac{d^{k_i}}{dt_i^{k_i}} t_i^{\ell_i} \quad (22)$$

$$= \prod_{i=1}^n I\{k_i \leq \ell_i\} \frac{\ell_i!}{(\ell_i - k_i)!} t_i^{\ell_i - k_i} \quad (23)$$

Thus,

$$\frac{d^{\sum_{i=1}^n k_i}}{d^{k_1} d^{k_2} \dots d^{k_n}} \left(\sum_{ij} \sigma_{ij} t_i t_j \right)^\ell = \sum_{\{(\ell_1, \dots, \ell_n) | \sum_i \ell_i = 2\ell\}} \sum_{\{(\ell_{11}, \ell_{12}, \dots, \ell_{nn}) | \sum_i (\ell_{ih} + \ell_{hi}) = \ell_h, h=1, \dots, n\}} \left(\binom{\ell}{\ell_{11} \dots \ell_{nn}} \prod_{ij} \sigma_{ij}^{\ell_{ij}} \right) \prod_{i=1}^n I\{k_i \leq \ell_i\} \frac{\ell_i!}{(\ell_i - k_i)!} t_i^{\ell_i - k_i} \quad (24)$$

Incorporating the constants from Equation 11, noting again that the negative signs will cancel, the full sum is

$$\sum_{\ell=0}^{\infty} \left(\frac{1}{2}\right)^{\ell} / \ell! \sum_{\{(\ell_1, \dots, \ell_n) | \sum_i \ell_i = 2\ell\}} \left(\sum_{\{(\ell_{11}, \ell_{12}, \dots, \ell_{nn}) | \sum_j (\ell_{hj} + \ell_{jh}) = \ell_i, h=1, \dots, n\}} \binom{\ell}{\ell_{11} \dots \ell_{nn}} \prod_{ij} \sigma_{ij}^{\ell_{ij}} \right) \prod_{i=1}^n I\{k_i \leq \ell_i\} \frac{\ell_i!}{(\ell_i - k_i)!} t_i^{\ell_i - k_i} \quad (25)$$

Setting $t = 0$, only terms with $\ell_i = k_i$ for all i will remain. Otherwise, the only ℓ in the infinite sum which occurs is for $\ell = \sum_{i=1}^n k_i/2$. So this reduces to

$$\left(\frac{1}{2}\right)^{\sum_{i=1}^n k_i/2} / \left(\sum_{i=1}^n k_i/2\right)! \left(\sum_{\{(\ell_{11}, \dots, \ell_{nn}) | \sum_j (\ell_{hj} + \ell_{jh}) = k_i, h=1, \dots, n\}} \binom{\sum_{i=1}^n k_i/2}{\ell_{11} \dots \ell_{nn}} \prod_{ij} \sigma_{ij}^{\ell_{ij}} \right) \prod_{i=1}^n k_i! \quad (26)$$

Rearranging the terms, we have

$$m_{k_1, \dots, k_n} = C \sum_{\{(\ell_{11}, \ell_{12}, \dots, \ell_{nn}) | \sum_{j=1}^n (\ell_{jh} + \ell_{hj}) = k_h, h=1, \dots, n\}} \binom{\sum_{i=1}^n k_i/2}{\ell_{11} \dots \ell_{nn}} \prod_{ij} \sigma_{ij}^{\ell_{ij}} \quad (27)$$

where

$$C = \frac{1}{2^{\sum_{i=1}^n k_i/2}} \left(\prod_{i=1}^n k_i! \right) / \left(\sum_{i=1}^n k_i/2 \right)! \quad (28)$$

This formula shows that evaluating m_{k_1, \dots, k_n} symbolically requires enumerating all $n \times n$ -dimensional matrices of non-negative integers, (ℓ_{ij}) , which satisfy the condition

$$\sum_{j=1}^n (\ell_{ji} + \ell_{ij}) = k_i, \quad i = 1, \dots, n \quad (29)$$

Conceding now that the symbols σ_{ij} and σ_{ji} signify the same entity, we can search for (ℓ_{ij}) by looking only at terms $\prod_{ij} \sigma_{ij}^{\ell_{ij}}$ for which $i \leq j$. In fact, any other matrix for the term can be obtained by decrementing ℓ_{ij} and incrementing ℓ_{ji} by the same integer for one or more subscripts for which $i < j$. For any $\sigma_{ij}^{\ell_{ij}}$ in the term, this can be done in $\ell_{ij} + 1$ ways. So there are a total of $\prod_{i < j} (\ell_{ij} + 1)$ transpositions. The multinomial coefficients derived above must be applied separately to each of these (ℓ_{ij}) matrices. Thus, the full coefficient for a matrix will include as a multiplier the sum of these coefficients over all of the $\prod_{i < j} (\ell_{ij} + 1)$ transposed matrices. Let Υ be the set of upper-triangular integer matrices, and, for any $(\ell_{ij}) \in \Upsilon$, let $\Lambda((\ell_{ij}))$ be the set of all integer matrices (h_{ij}) obtained by so transposing (ℓ_{ij}) . Then the sum above can be decomposed in terms of Υ and $\Lambda((\ell_{ij}))$ for each $(\ell_{ij}) \in \Upsilon$:

$$m_{k_1, \dots, k_n} = C \sum_{\{(\ell_{11}, \ell_{12}, \dots, \ell_{nn}) \in \Upsilon | \sum_{j=1}^n (\ell_{jh} + \ell_{hj}) = k_h, h=1, \dots, n\}} \sum_{\{(h_{ij}) \in \Lambda((\ell_{ij}))\}} \binom{\sum_{i=1}^n k_i/2}{h_{11} \dots h_{nn}} \prod_{ij} \sigma_{ij}^{h_{ij}} \quad (30)$$

But by symmetry, the products in (σ_{ij}) are the same for each member of $\Lambda((\ell_{ij}))$, specifically $\prod_{ij} \sigma_{ij}^{\ell_{ij}}$. So the final formula is

$$m_{k_1, \dots, k_n} = \left[\frac{1}{2}^{\sum_{i=1}^n k_i/2} \left(\prod_{i=1}^n k_i! \right) / \left(\sum_{i=1}^n k_i/2 \right)! \right] \sum_{\{(\ell_{11}, \ell_{12}, \dots, \ell_{nn}) \in \mathcal{T} | \sum_{j=1}^n (\ell_{jh} + \ell_{hj}) = k_h, h=1, \dots, n\}} \left[\sum_{\{(h_{ij}) \in \Lambda((\ell_{ij}))\}} \left(\frac{\sum_{i=1}^n k_i/2}{h_{11} \dots h_{nn}} \right) \right] \prod_{ij} \sigma_{ij}^{\ell_{ij}} \quad (31)$$

3. Discussion

Formula 31 was implemented in R (R Development Core Team 2009) with a recursive function that determines the set of upper-triangular integer matrices that satisfy Criterion 29. A second function calculates their associated coefficients. Additional functions were written to create L^AT_EX (L^AT_EX3 Project Team 2009) code to display the moments symbolically, and to calculate the moments for specified variance-covariance matrices.

The potential for complexity in these computations is seen from the results in Table 1. In this table, n is the dimension of the multivariate vector and $\#(\sigma_{ij})$ is the number of distinct elements in the variance-covariance matrix, $N = \frac{n(n+1)}{2}$. *Size* is measured by two values, M and r . M is the total of the exponents of the moment, $M = \sum_{i=1}^n k_i$. The value r is the number of terms for a moment $E[X_1^1, \dots, X_n^1]$ with all exponents equal to 1, which is $(2M-1)!/(2^{M-1}(M-1)!)$ (Wikipedia 2009). *Example* is a moment of the given *Size*. *Potential Terms* is a maximum for the number of (ℓ_{ij}) matrices to be checked for this example, determined as the product of $1 + \max(k_i, k_j)$ over $i \neq j$ times the product of $1 + [k_i/2]$ over i , where $[]$ denotes truncation. The last column, *# Terms*, is the actual number of terms in the moment as determined by the functions.

It is clear that computation of high-order moments will be very intensive. Kan reports similar computational difficulties. More efficient or targeted algorithms for searching the matrices in Criterion 29 would allow higher order moments to be computed. However, it is likely that the problem is intractable as described by Garey and Johnson (1979). For example, symbolic computation of the central moments $E[X_1^{k_1} \dots X_n^{k_n}]$ for $(k_1, \dots, k_n) < (k, \dots, k)$ for a fixed k may be NP-hard in n , or computation of $E[X_1^{k_1} \dots X_n^{k_n}]$ may be NP-hard in $\max(k_1, \dots, k_n)$ for fixed n .

The formula derived here could be expanded to incorporate mean terms, which would allow computation of non-central moments. These moments could also be used to approximate other integrals integrated against the multivariate normal distribution by using a Taylor expansion in several variables (Fulks 1961). Such an approximation would be a linear combination of non-central moments.

Finally, Criterion 29 might arise in other contexts, such as networks (Stergiou and Siganos 1996). For example, suppose that there are n airports and airport i can accommodate k_i arrivals or departures on a day, where a plane may take off and land at the same airport. This network is illustrated in Figure 1. The problem is to determine the set of flights between airports that totally expend the capacities, k_i , of all airports. For this problem, ℓ_{ij} represents

n	$\#(\sigma_{ij})$	Size		Example	Potential Terms	# Terms
	N	M	r	$E[X_1^{k_1} \dots X_n^{k_n}]$		
2	3	2	1	$E[X_1^1 X_2^1]$	2	1
2	3	6	15	$E[X_1^3 X_2^3]$	16	2
2	3	20	654,729,075	$E[X_1^{10} X_2^{10}]$	396	6
4	10	8	105	$E[X_1^2 X_2^2 X_3^2 X_4^2]$	11,664	17
4	10	12	10,395	$E[X_1^1 X_2^3 X_3^4 X_4^4]$	225,000	27
4	10	20	654,729,075	$E[X_1^5 X_2^5 X_3^5 X_4^5]$	3,779,136	306
6	21	12	10,395	$E[X_1^2 X_2^2 X_3^2 X_4^2 X_5^2 X_6^2]$	918,330,048	388
6	21	18	34,459,425	$E[X_1^1 X_2^2 X_3^3 X_4^4 X_5^4 X_6^4]$	1.265625×10^{12}	2,082
8	36	16	2,027,025	$E[X_1^2 X_2^2 X_3^2 X_4^2 X_5^2 X_6^2 X_7^2 X_8^2]$	5.856459×10^{15}	18,155
8	36	24	316,234,143,225	$E[X_1^3 X_2^3 X_3^3 X_4^3 X_5^3 X_6^3 X_7^3 X_8^3]$	1.844674×10^{19}	$\geq 287,800^*$
9	45	18	34,459,425	$E[X_1^2 X_2^2 X_3^2 X_4^2 X_5^2 X_6^2 X_7^2 X_8^2 X_9^2]$	7.68484×10^{19}	6,763,895

Table 1: Examples of sizes of moment problem (*: The function aborted after 2 days due a space allocation problem.)

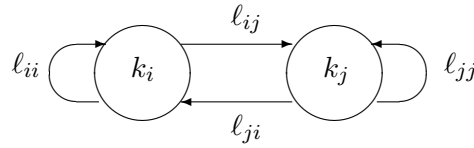


Figure 1: Airport capacity network.

the number of flights from airport i to airport j , and ℓ_{ii} is the number of flights which start and end at airport i . The set of flights is the set satisfying Criterion 29.

4. R functions to compute normal multivariate moments

In the R functions to compute the central moments, upper-triangular matrices (ℓ_{ij}) for n dimensions are represented as vectors of length $n(n+1)$, with row 1 followed by row 2, etc. For example, for $n = 2$, (ℓ_{ij}) is represented as $(\ell_{11}, \ell_{12}, \ell_{22})$. Each such matrix represents the exponents for a single product of σ_{ij} s. For example, $(1, 2, 0)$ represents $\sigma_{11}^1 \sigma_{12}^2 \sigma_{22}^0 = \sigma_{11} \sigma_{12}^2$. The representations are accumulated and stored internally, raising the possibility of space allocation problems as encountered in the second to last example in Table 1. This problem could be alleviated by saving the representation matrix to a file instead.

The function `multmoments` searches the integer matrices for those satisfying Criterion 29. This is a recursive function which implements a branch-and-bound algorithm. The function `multmoments` is called by `callmultmoments`. This function initializes variables, determines the coefficients of the terms from the upper-triangular representations, and returns a list consisting of the original moment vector, the set of representations, and the corresponding set of coefficients. This list is set to class `moment`. The `moment` class has four methods: `print`, `toLatex`, `evaluate`, and `simulate`.

The `print` method prints a moment object, usually the output of `callmultmoments`, showing a mathematical representation of the moment, followed by the rows of the representation with the corresponding coefficient attached.

The `toLatex` method uses a moment object, usually the output of `callmultmoments`, to determine the \LaTeX code for the moment sorted lexicographically. Note that it inserts double backslashes where \LaTeX requires a backslash; these can be reset to single backslashes by writing the output to a file using the R base function `writeLines`, as illustrated below.

The `evaluate` method determines the value of a `moment` object for a specified variance-covariance matrix Σ , which must be represented as an upper-triangular matrix in vector form.

The `simulate` method uses Monte Carlo integration Rizzo (2008) to numerically approximate a `moment` object for a specified mean and variance-covariance matrix Σ (represented as a square matrix in vector form), with a specified number of random samples. Note that `simulate` uses only the moment definition, not the representation, so can be used with any moment in vector notation by converting the vector to a moment object. The `simulate` method uses the `rmvnorm` function from the `mvtnorm` package (Genz et al. 2009).

Computation of the moments was validated by comparison to specific published cases and to known types such as $E[X^k]$ and $E[X_1^1, \dots, X_n^1]$, through consistency checks, and to comparison to numerically-computed moments. For consistency, the representations for moments which

are permutations of each other must be the same within ordering across rows and columns; for example, $E[X_1^2 X_2^4]$ and $E[X_1^4 X_2^2]$ have the same representations within ordering. Also, a moment containing one or more odd powers will evaluate to 0 for a specified covariance matrix if the component of the random variable corresponding to one of the odd powers is independent of the other components, that is, if the off-diagonal covariance terms are zero for this component. The functions were confirmed to have this property for a number of cases.

Further validation was done using Monte Carlo integration. Various moments of dimension up to six were compared for two values of the covariance matrix, the identity matrix and a covariance matrix computed from ten random normal vectors obtained using the identity covariance matrix and then adding 1 to the diagonal elements to increase the variability. Forty estimates of the moment were then computed using the `simulate` method, each obtained from one million randomly generated multivariate normal vectors. From these estimates, 95% confidence intervals were computed and compared to the moment estimates using the `estimate` method. In the 37 experiments done using an identity matrix, the moments computed from the symbolic representations fell outside the confidence interval in three (8.1%) of the cases. However, in all three cases `callmultmoments` produced the value of zero, which is correct since the moments contained odd powers and the components of the vectors were independent. In the 37 experiments done using a randomly-generated matrix, the moments computed from the symbolic representations fell outside the confidence interval in two (5.4%) of the cases ($E[X_1^2 X_2^9 X_3^{11}]$ and $E[X_1^2 X_2^4 X_3^7 X_4^7 X_5^8]$), where the symbolically computed values lay slightly outside the confidence intervals. These two cases were run again using five million random normal vectors for each of the 40 estimates, and the moments computed from the symbolic representations fell inside the confidence intervals. These results provide further evidence that the functions given here are correct, and in fact are superior to numerical integration in obtaining moments, since it only requires a simple evaluation of a polynomial with integer coefficients. Numerical integration using adaptive methods was also implemented but worked poorly for higher dimensions (Kuonen 2003).

5. Examples of computing central moments

The following moment is computed using the code given below:

$$\begin{aligned}
 E[X_1^1 X_2^2 X_3^3 X_4^4] = & \\
 & 18\sigma_{1,2}\sigma_{2,3}\sigma_{3,3}\sigma_{4,4}^2 + 72\sigma_{1,2}\sigma_{2,3}\sigma_{3,4}^2\sigma_{4,4} + 72\sigma_{1,2}\sigma_{2,4}\sigma_{3,3}\sigma_{3,4}\sigma_{4,4} + 48\sigma_{1,2}\sigma_{2,4}\sigma_{3,4}^3 + \\
 & 9\sigma_{1,3}\sigma_{2,2}\sigma_{3,3}\sigma_{4,4}^2 + 36\sigma_{1,3}\sigma_{2,2}\sigma_{3,4}^2\sigma_{4,4} + 18\sigma_{1,3}\sigma_{2,3}^2\sigma_{4,4}^2 + 144\sigma_{1,3}\sigma_{2,3}\sigma_{2,4}\sigma_{3,4}\sigma_{4,4} + \\
 & 36\sigma_{1,3}\sigma_{2,4}^2\sigma_{3,3}\sigma_{4,4} + 72\sigma_{1,3}\sigma_{2,4}^2\sigma_{3,4}^2 + 36\sigma_{1,4}\sigma_{2,2}\sigma_{3,3}\sigma_{3,4}\sigma_{4,4} + 24\sigma_{1,4}\sigma_{2,2}\sigma_{3,4}^3 + \\
 & 72\sigma_{1,4}\sigma_{2,3}^2\sigma_{3,4}\sigma_{4,4} + 72\sigma_{1,4}\sigma_{2,3}\sigma_{2,4}\sigma_{3,3}\sigma_{4,4} + 144\sigma_{1,4}\sigma_{2,3}\sigma_{2,4}\sigma_{3,4}^2 + 72\sigma_{1,4}\sigma_{2,4}^2\sigma_{3,3}\sigma_{3,4}
 \end{aligned} \tag{32}$$

The use of the `toLatex` and `evaluate` methods and `writeLines` is also illustrated. The file created by `writeLines` can be included in a \LaTeX document using the `\input` command, or can be included in Sweave as done here.

Compute the representation of a central moment (callmultmoments)

The following code calculates a central moment and shows the three components, `moment`,

representation, and coefficients.

```
R> m1234 <- callmultmoments(c(1,2,3,4))
R> unclass(m1234)
```

\$moment

```
[1] 1 2 3 4
```

\$representation

	S(1,1)	S(1,2)	S(1,3)	S(1,4)	S(2,2)	S(2,3)	S(2,4)	S(3,3)	S(3,4)	S(4,4)
1	0	0	0	1	0	0	2	1	1	0
2	0	0	0	1	0	1	1	0	2	0
3	0	0	0	1	0	1	1	1	0	1
4	0	0	0	1	0	2	0	0	1	1
5	0	0	0	1	1	0	0	0	3	0
6	0	0	0	1	1	0	0	1	1	1
7	0	0	1	0	0	0	2	0	2	0
8	0	0	1	0	0	0	2	1	0	1
9	0	0	1	0	0	1	1	0	1	1
10	0	0	1	0	0	2	0	0	0	2
11	0	0	1	0	1	0	0	0	2	1
12	0	0	1	0	1	0	0	1	0	2
13	0	1	0	0	0	0	1	0	3	0
14	0	1	0	0	0	0	1	1	1	1
15	0	1	0	0	0	1	0	0	2	1
16	0	1	0	0	0	1	0	1	0	2

\$coefficients

rep 1	rep 2	rep 3	rep 4	rep 5	rep 6	rep 7	rep 8	rep 9	rep 10	rep 11
72	144	72	72	24	36	72	36	144	18	36
rep 12	rep 13	rep 14	rep 15	rep 16						
9	48	72	72	18						

Print a representation of a central moment (print method)

The following shows the result of using the `print` method with the moment in Equation 32.

```
R> m1234
```

E[$X_1^1 X_2^2 X_3^3 X_4^4$]:

	coef	S(1,1)	S(1,2)	S(1,3)	S(1,4)	S(2,2)	S(2,3)	S(2,4)	S(3,3)	S(3,4)	S(4,4)
1	72	0	0	0	1	0	0	2	1	1	0
2	144	0	0	0	1	0	1	1	0	2	0
3	72	0	0	0	1	0	1	1	1	0	1
4	72	0	0	0	1	0	2	0	0	1	1
5	24	0	0	0	1	1	0	0	0	3	0

6	36	0	0	0	1	1	0	0	1	1	1
7	72	0	0	1	0	0	0	2	0	2	0
8	36	0	0	1	0	0	0	2	1	0	1
9	144	0	0	1	0	0	1	1	0	1	1
10	18	0	0	1	0	0	2	0	0	0	2
11	36	0	0	1	0	1	0	0	0	2	1
12	9	0	0	1	0	1	0	0	1	0	2
13	48	0	1	0	0	0	0	1	0	3	0
14	72	0	1	0	0	0	0	1	1	1	1
15	72	0	1	0	0	0	1	0	0	2	1
16	18	0	1	0	0	0	1	0	1	0	2

Compute the \LaTeX representation of a central moment (`toLatex` method)

The following shows the computation of the representation of the central moment in Equation 32. (The output was slightly edited for display in this manuscript.)

```
R> toLatex(m1234)
```

```
[1] "E[ X_{ 1 }^{ 1 } X_{ 2 }^{ 2 } X_{ 3 }^{ 3 } X_{ 4 }^{ 4 } ] = \\\\"
[2] "18 \\sigma_{ 1 , 2 } \\sigma_{ 2 , 3 } \\sigma_{ 3 , 3 } \\sigma_{ 4 , 4 }^{ 2 } + \"
[3] "72 \\sigma_{ 1 , 2 } \\sigma_{ 2 , 3 } \\sigma_{ 3 , 4 }^{ 2 } \\sigma_{ 4 , 4 } + \"
[4] "72 \\sigma_{ 1 , 2 } \\sigma_{ 2 , 4 } \\sigma_{ 3 , 3 } \\sigma_{ 3 , 4 } \\sigma_{ 4 , 4 } + \"
[5] "48 \\sigma_{ 1 , 2 } \\sigma_{ 2 , 4 } \\sigma_{ 3 , 4 }^{ 3 } + \\\\"
[6] "9 \\sigma_{ 1 , 3 } \\sigma_{ 2 , 2 } \\sigma_{ 3 , 3 } \\sigma_{ 4 , 4 }^{ 2 } + \"
[7] "36 \\sigma_{ 1 , 3 } \\sigma_{ 2 , 2 } \\sigma_{ 3 , 4 }^{ 2 } \\sigma_{ 4 , 4 } + \"
[8] "18 \\sigma_{ 1 , 3 } \\sigma_{ 2 , 3 }^{ 2 } \\sigma_{ 4 , 4 }^{ 2 } + \"
[9] "144 \\sigma_{ 1 , 3 } \\sigma_{ 2 , 3 } \\sigma_{ 2 , 4 } \\sigma_{ 3 , 4 } \\sigma_{ 4 , 4 } + \\\\"
[10] "36 \\sigma_{ 1 , 3 } \\sigma_{ 2 , 4 }^{ 2 } \\sigma_{ 3 , 3 } \\sigma_{ 4 , 4 } + \"
[11] "72 \\sigma_{ 1 , 3 } \\sigma_{ 2 , 4 }^{ 2 } \\sigma_{ 3 , 4 }^{ 2 } + \"
[12] "36 \\sigma_{ 1 , 4 } \\sigma_{ 2 , 2 } \\sigma_{ 3 , 3 } \\sigma_{ 3 , 4 } \\sigma_{ 4 , 4 } + \"
[13] "24 \\sigma_{ 1 , 4 } \\sigma_{ 2 , 2 } \\sigma_{ 3 , 4 }^{ 3 } + \\\\"
[14] "72 \\sigma_{ 1 , 4 } \\sigma_{ 2 , 3 }^{ 2 } \\sigma_{ 3 , 4 } \\sigma_{ 4 , 4 } + \"
```

```

[15] "72 \\sigma_{ 1 , 4 } \\sigma_{ 2 , 3 } \\sigma_{ 2 , 4 }
      \\sigma_{ 3 , 3 } \\sigma_{ 4 , 4 } + "
[16] "144 \\sigma_{ 1 , 4 } \\sigma_{ 2 , 3 } \\sigma_{ 2 , 4 }
      \\sigma_{ 3 , 4 }^{ 2 } + "
[17] "72 \\sigma_{ 1 , 4 } \\sigma_{ 2 , 4 }^{ 2 } \\sigma_{ 3 , 3 }
      \\sigma_{ 3 , 4 } \\\\"

```

The L^AT_EX representation can be written to a file using the `writeLines` function as follows:

```
R> writeLines(toLatex(m1234), "yourfilename")
```

Compute a value of a central moment (evaluate method)

The code below evaluates the moment at the following variance-covariance matrix:

	[,1]	[,2]	[,3]	[,4]
[1,]	4	2	1	1
[2,]	2	3	1	1
[3,]	1	1	2	1
[4,]	1	1	1	2

```
R> evaluate(m1234, c(4, 2, 1, 1, 3, 1, 1, 2, 1, 2))
```

```
[1] 3480
```

Estimate a central moment using simulation (simulate method)

The value of $E[X_1^1 X_2^2 X_3^3 X_4^4]$ when X has a normal distribution with mean $\mu = (1, 2, 0, 3)$ and the same variance-covariance matrix as above could be estimated using `simulate` with 1000 random samples:

```
R> simulate(m1234, 1000, NULL, c(1, 2, 0, 3),
+       c(4, 2, 1, 1, 2, 3, 1, 1, 1, 1, 2, 1, 1, 1, 1, 2))
```

```
[1] 73598.17
```

Acknowledgments

The author would like to thank Dr. Robin Hankin and Dr. Achim Zeileis for several suggestions and assistance in constructing the **symmoments** package.

References

- Anderson TW (1971). *An Introduction to Multivariate Statistical Analysis*. John Wiley & Sons, New York.
- Fulks W (1961). *Advanced Calculus*. John Wiley & Sons, New York.
- Garey MR, Johnson DS (1979). *Computers and Intractability, A Guide to the Theory of NP-Completeness*. W. H. Freeman and Company, New York.
- Genz A, Bretz F, Miwa T, Mi X, Leisch F, Scheipl F, Hothorn T (2009). *mvtnorm: Multivariate Normal and t Distributions*. R package version 0.9-8, URL <http://CRAN.R-project.org/package=mvtnorm>.
- Isserlis L (1918). “On a Formula for the Product-Moment Coefficient of Any Order of a Normal Frequency Distribution in Any Number of Variables.” *Biometrika*, **12**(1–2), 134–139.
- Kan R (2008). “From Moments of Sums to Moments of Product.” *Journal of Multivariate Analysis*, **99**, 542–554.
- Kuonen D (2003). “Numerical Integration in S-PLUS or R: A Survey.” *Journal of Statistical Software*, **8**(13), 1–14. URL <http://www.jstatsoft.org/v08/i13>.
- L^AT_EX3 Project Team (2009). *L^AT_EX – A Document Preparation System*. The L^AT_EX3 Project. URL <http://www.LaTeX-project.org/>.
- Muirhead RJ (1982). *Aspects of Multivariate Statistical Theory*. John Wiley & Sons, New York.
- R Development Core Team (2009). *R: A Language and Environment for Statistical Computing*. R Foundation for Statistical Computing, Vienna, Austria. ISBN 3-900051-07-0, URL <http://www.R-project.org/>.
- Rizzo ML (2008). *Statistical Computing with R*. Chapman & Hall/CRC, London.
- Stergiou C, Siganos D (1996). “Neural Networks.” *Technical report*, Department of Computing, Imperial College, Surveys and Presentations in Information Systems Engineering (SURPRISE). URL http://www.doc.ic.ac.uk/~nd/surprise_96/journal/vol4/cs11/report.html.
- Tracy DS, Sultan SA (1993). “Higher Order Moments of Multivariate Normal Distribution Using Matrix Derivatives.” *Stochastic Analysis and Applications*, **11**, 337–348.
- Wikipedia (2009). *Multivariate Normal Distribution* — *Wikipedia, The Free Encyclopedia*. Last revision from 2009-12-08 17:01 UTC, URL http://en.wikipedia.org/w/index.php?title=Multivariate_normal_distribution&oldid=330485779.

Affiliation:

Kem Phillips
314 Rileyville Road
Ringo
NJ, 08551, United States of America
E-mail: kemphillips@comcast.net